# Including prior knowledge in machine learning for genomic data

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Mines ParisTech / Curie Institute / Inserm

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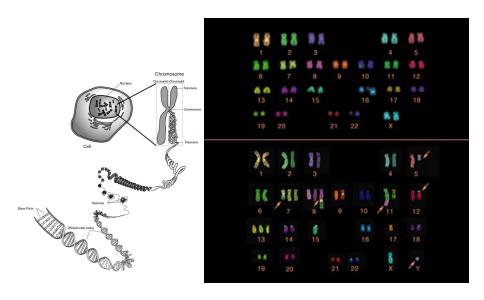
### Outline

- Motivations
- Pinding multiple change-points in a single profile
- Finding multiple change-points shared by many signals
- Supervised classification of genomic profiles
- 5 Learning molecular classifiers with network information
- 6 Conclusion

#### **Outline**

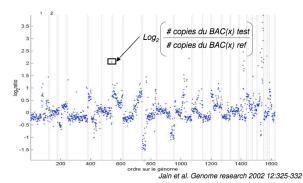
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#### Chromosomic aberrations in cancer

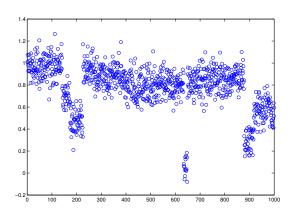


## Comparative Genomic Hybridization (CGH)

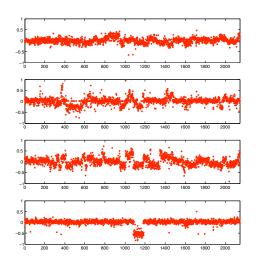




# Can we identify breakpoints and "smooth" each profile?

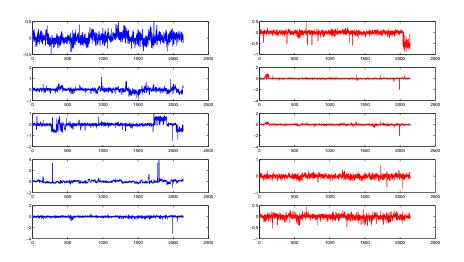


## Can we detect frequent breakpoints?



A collection of bladder tumour copy number profiles.

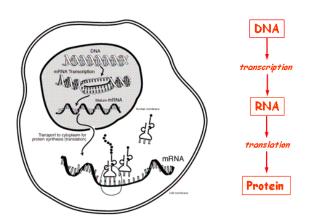
## Can we detect discriminative patterns?



Aggressive (left) vs non-aggressive (right) melanoma.

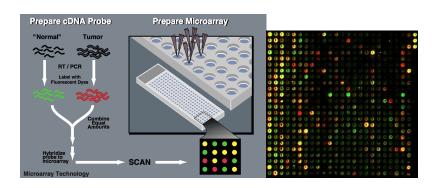
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## DNA → RNA → protein



- CGH shows the (static) DNA
- Cancer cells have also abnormal (dynamic) gene expression (= transcription)

## Tissue profiling with DNA chips



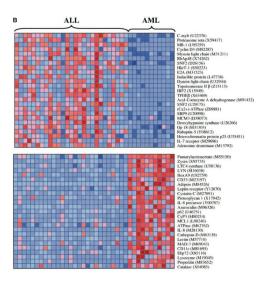
#### Data

- Gene expression measures for more than 10k genes
- Measured typically on less than 100 samples of two (or more) different classes (e.g., different tumors)

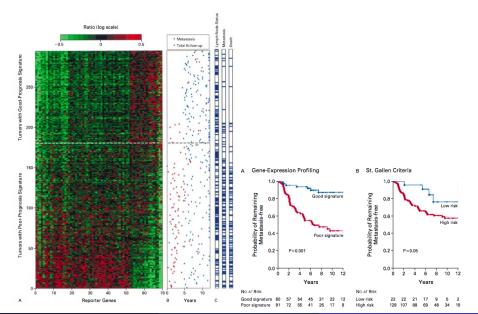
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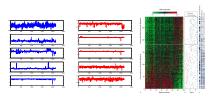
## Can we identify the cancer subtype? (diagnosis)



## Can we predict the future evolution? (prognosis)



## Summary



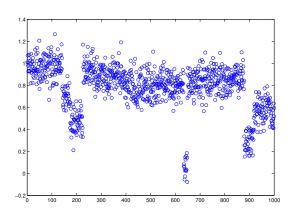
- Many problems...
- Data are high-dimensional, but "structured"
- Classification accuracy is not all, interpretation is necessary (pattern discovery)
- A general strategy

$$\min R(\beta) + \lambda \Omega(\beta)$$

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## The problem

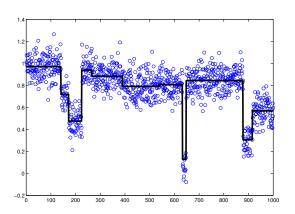


- Let  $Y \in \mathbb{R}^p$  the signal
- We want to find a piecewise constant approximation  $\hat{U} \in \mathbb{R}^p$  with at most k change-points.

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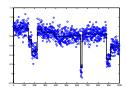
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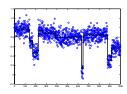
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• We can define an "optimal" piecewise constant approximation  $\hat{U} \in \mathbb{R}^p$  as the solution of

$$\min_{U\in\mathbb{R}^p}\|Y-U\|^2$$
 such that  $\sum_{i=1}^{p-1}\mathbf{1}\left(U_{i+1}\neq U_i\right)\leq k$ 

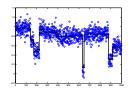
- ullet This is an optimization problem over the  $\binom{
  ho}{k}$  partitions
- Dynamic programming finds the solution in  $O(p^2k)$  in time and  $O(p^2)$  in memory
- But: does not scale to  $p = 10^6 \sim 10^9$ .



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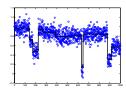
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## Promoting sparsity with the $\ell_1$ penalty

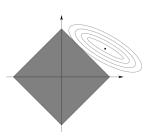
#### The $\ell_1$ penalty (Tibshirani, 1996; Chen et al., 1998)

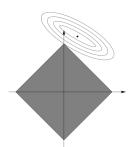
If  $R(\beta)$  is convex and "smooth", the solution of

$$\min_{\beta \in \mathbb{R}^p} R(\beta) + \lambda \sum_{i=1}^p |\beta_i|$$

is usually sparse.

Geometric interpretation with p=2





## Promoting piecewise constant profiles penalty

#### The total variation / variable fusion penalty

If  $R(\beta)$  is convex and "smooth", the solution of

$$\min_{\beta \in \mathbb{R}^p} R(\beta) + \lambda \sum_{i=1}^{p-1} |\beta_{i+1} - \beta_i|$$

is usually piecewise constant (Rudin et al., 1992; Land and Friedman, 1996).

#### Proof:

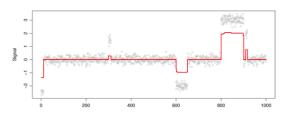
- Change of variable  $u_i = \beta_{i+1} \beta_i$ ,  $u_0 = \beta_1$
- We obtain a Lasso problem in  $u \in \mathbb{R}^{p-1}$
- u sparse means  $\beta$  piecewise constant

## TV signal approximator

$$\min_{\beta \in \mathbb{R}^p} \| Y - \beta \|^2 \quad \text{such that} \quad \sum_{i=1}^{p-1} |\beta_{i+1} - \beta_i| \le \mu$$

Adding additional constraints does not change the change-points:

- $\sum_{i=1}^{p} |\beta_i| \le \nu$  (Tibshirani et al., 2005; Tibshirani and Wang, 2008)
- $\sum_{i=1}^{p} \beta_i^2 \le \nu$  (Mairal et al. 2010)



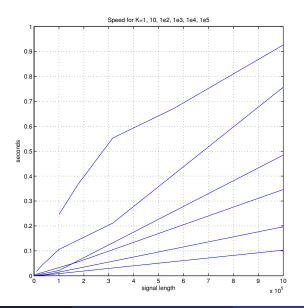
## Solving TV signal approximator

$$\min_{\beta \in \mathbb{R}^p} \| Y - \beta \|^2 \quad \text{such that} \quad \sum_{i=1}^{p-1} |\beta_{i+1} - \beta_i| \le \mu$$

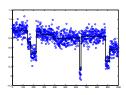
- QP with sparse linear constraints in  $O(p^2)$  -> 135 min for  $p = 10^5$  (Tibshirani and Wang, 2008)
- Coordinate descent-like method O(p)? -> 3s s for  $p = 10^5$  (Friedman et al., 2007)
- For all  $\mu$  with the LARS in O(pK) (Harchaoui and Levy-Leduc, 2008)
- For all  $\mu$  in  $O(p \ln p)$  (Hoefling, 2009)
- For the first K change-points in  $O(p \ln K)$  (Bleakley and V., 2010)

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# Speed trial : 2 s. for K = 100, $p = 10^7$



## Summary

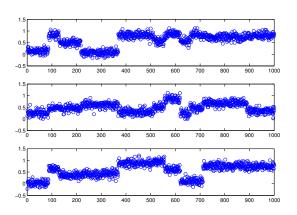


- A fast method for multiple change-point detection
- An embedded method that boils down to a dichotomic wrapper method (very different from dynamic programming)

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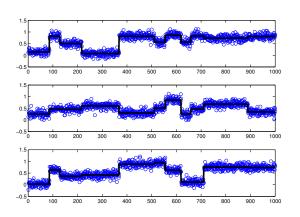


- Let  $Y \in \mathbb{R}^{p \times n}$  the *n* signals of length *p*
- We want to find a piecewise constant approximation  $\hat{U} \in \mathbb{R}^{p \times n}$  with at most k change-points.

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## The problem

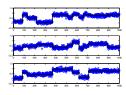


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## "Optimal" segmentation by dynamic programming



• Define the "optimal" piecewise constant approximation  $\hat{U} \in \mathbb{R}^{p \times n}$  of Y as the solution of

$$\min_{U \in \mathbb{R}^{p imes n}} \parallel Y - U \parallel^2 \quad \text{such that} \quad \sum_{i=1}^{p-1} \mathbf{1} \left( U_{i+1,ullet} 
eq U_{i,ullet} 
ight) \leq k$$

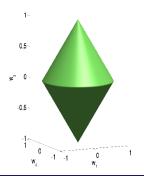
- DP finds the solution in  $O(p^2kn)$  in time and  $O(p^2)$  in memory
- But: does not scale to  $p = 10^6 \sim 10^9...$

## Selecting pre-defined groups of variables

#### Group lasso (Yuan & Lin, 2006)

If groups of covariates are likely to be selected together, the  $\ell_1/\ell_2$ -norm induces sparse solutions at the group level:

$$\Omega_{group}(w) = \sum_{g} \|w_{g}\|_{2}$$



$$\Omega(w_1, w_2, w_3) = \|(w_1, w_2)\|_2 + \|w_3\|_2$$
$$= \sqrt{w_1^2 + w_2^2} + \sqrt{w_3^2}$$

## TV approximator for many signals

Replace

$$\min_{U \in \mathbb{R}^{p \times n}} \| Y - U \|^2$$
 such that  $\sum_{i=1}^{p-1} \mathbf{1} \left( U_{i+1,\bullet} \neq U_{i,\bullet} \right) \leq k$ 

by

$$\min_{U\in\mathbb{R}^{\rho imes n}}\|Y-U\|^2$$
 such that  $\sum_{i=1}^{\rho-1}w_i\|U_{i+1,ullet}-U_{i,ullet}\|\leq \mu$ 

#### Questions

- Practice: can we solve it efficiently?
- Theory: does it benefit from increasing *p* (for *n* fixed)?

## TV approximator as a group Lasso problem

Make the change of variables:

$$\gamma = U_{1,\bullet}$$
,  
 $\beta_{i,\bullet} = w_i \left( U_{i+1,\bullet} - U_{i,\bullet} \right)$  for  $i = 1, \dots, p-1$ .

 TV approximator is then equivalent to the following group Lasso problem (Yuan and Lin, 2006):

$$\min_{\beta \in \mathbb{R}^{(p-1) \times n}} \| \bar{Y} - \bar{X}\beta \|^2 + \lambda \sum_{i=1}^{p-1} \| \beta_{i,\bullet} \|,$$

where  $\bar{Y}$  is the centered signal matrix and  $\bar{X}$  is a particular  $(p-1)\times(p-1)$  design matrix.

## TV approximator implementation

$$\min_{\beta \in \mathbb{R}^{(\rho-1) \times n}} \| \bar{Y} - \bar{X}\beta \|^2 + \lambda \sum_{i=1}^{\rho-1} \| \beta_{i,\bullet} \|,$$

#### **Theorem**

The TV approximator can be solved efficiently:

- approximately with the group LARS in O(npk) in time and O(np) in memory
- exactly with a block coordinate descent + active set method in O(np) in memory

## Proof: computational tricks...

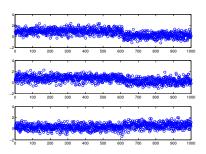
## Although $\bar{X}$ is $(p-1) \times (p-1)$ :

- For any  $R \in \mathbb{R}^{p \times n}$ , we can compute  $C = \bar{X}^T R$  in O(np) operations and memory
- For any two subset of indices  $A = (a_1, \ldots, a_{|A|})$  and  $B = (b_1, \ldots, b_{|B|})$  in [1, p-1], we can compute  $\bar{X}_{\bullet,A}^{\top} \bar{X}_{\bullet,B}$  in O(|A||B|) in time and memory
- For any  $A = (a_1, \ldots, a_{|A|})$ , set of distinct indices with  $1 \le a_1 < \ldots < a_{|A|} \le p-1$ , and for any  $|A| \times n$  matrix R, we can compute  $C = \left(\bar{X}_{\bullet,A}^{\top} \bar{X}_{\bullet,A}\right)^{-1} R$  in O(|A|n) in time and memory

## Consistency for a single change-point

Suppose a single change-point:

- at position  $u = \alpha p$
- with increments  $(\beta_i)_{i=1,\dots,n}$  s.t.  $\bar{\beta}^2 = \lim_{k\to\infty} \frac{1}{n} \sum_{i=1}^n \beta_i^2$
- ullet corrupted by i.i.d. Gaussian noise of variance  $\sigma^2$



Does the TV approximator correctly estimate the first change-point as *p* increases?

# Consistency of the unweighted TV approximator

$$\min_{U \in \mathbb{R}^{p \times n}} \| Y - U \|^2 \quad \text{such that} \quad \sum_{i=1}^{p-1} \| U_{i+1,\bullet} - U_{i,\bullet} \| \le \mu$$

#### **Theorem**

The unweighted TV approximator finds the correct change-point with probability tending to 1 (resp. 0) as  $n \to +\infty$  if  $\sigma^2 < \tilde{\sigma}_{\alpha}^2$  (resp.  $\sigma^2 > \tilde{\sigma}_{\alpha}^2$ ), where

$$\tilde{\sigma}_{\alpha}^{2} = p\bar{\beta}^{2} \frac{(1-\alpha)^{2}(\alpha-\frac{1}{2p})}{\alpha-\frac{1}{2}-\frac{1}{2p}}.$$

- correct estimation on  $[p\epsilon, p(1-\epsilon)]$  with  $\epsilon = \sqrt{\frac{\sigma^2}{2p\bar{\beta}^2}} + o(p^{-1/2})$ .
- wrong estimation near the boundaries

# Consistency of the weighted TV approximator

$$\min_{\boldsymbol{U} \in \mathbb{R}^{p \times n}} \| \ \boldsymbol{Y} - \boldsymbol{U} \|^2 \quad \text{such that} \quad \sum_{i=1}^{p-1} \mathbf{\textit{w}}_i \| \boldsymbol{\textit{U}}_{i+1, \bullet} - \boldsymbol{\textit{U}}_{i, \bullet} \| \leq \mu$$

#### **Theorem**

The weighted TV approximator with weights

$$\forall i \in [1, p-1], \quad w_i = \sqrt{\frac{i(p-i)}{p}}$$

correctly finds the first change-point with probability tending to 1 as  $n \to +\infty$ .

- we see the benefit of increasing n
- we see the benefit of adding weights to the TV penalty

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## Proof sketch

• The first change-point  $\hat{i}$  found by TV approximator maximizes  $F_i = \|\hat{c}_{i,\bullet}\|^2$ , where

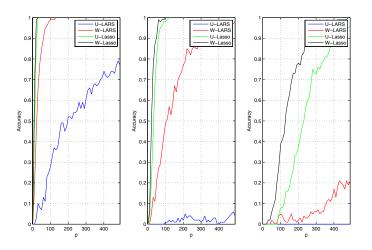
$$\hat{\mathbf{c}} = \bar{\mathbf{X}}^{\top} \bar{\mathbf{Y}} = \bar{\mathbf{X}}^{\top} \bar{\mathbf{X}} \beta^* + \bar{\mathbf{X}}^{\top} \mathbf{W}$$
.

•  $\hat{c}$  is Gaussian, and  $F_i$  is follows a non-central  $\chi^2$  distribution with

$$G_{i} = \frac{EF_{i}}{p} = \frac{i(p-i)}{pw_{i}^{2}}\sigma^{2} + \frac{\bar{\beta}^{2}}{w_{i}^{2}w_{u}^{2}p^{2}} \times \begin{cases} i^{2}\left(p-u\right)^{2} & \text{if } i \leq u\,, \\ u^{2}\left(p-i\right)^{2} & \text{otherwise}. \end{cases}$$

• We then just check when  $G_u = \max_i G_i$ 

# Consistent estimation of more change-points?



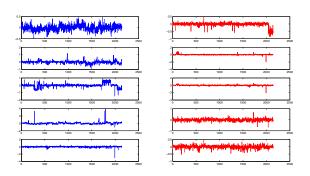
$$p = 100, k = 10, \bar{\beta}^2 = 1, \sigma^2 \in \{0.05; 0.2; 1\}$$

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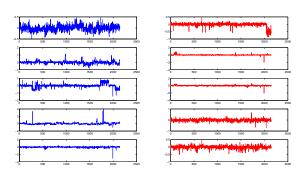
# The problem



- $x_1, \ldots, x_n \in \mathbb{R}^p$  the *n* profiles of length *p*
- $y_1, ..., y_n \in [-1, 1]$  the labels
- We want to learn a function  $f: \mathbb{R}^p \to [-1, 1]$

# Prior knowledge

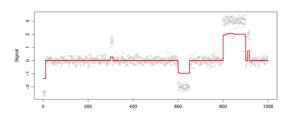
- Sparsity: not all positions should be discriminative, and we want to identify the predictive region (presence of oncogenes or tumor suppressor genes?)
- Piecewise constant: within a selected region, all probes should contribute equally



# Fused Lasso signal approximator (Tibshirani et al., 2005)

$$\min_{\beta \in \mathbb{R}^p} \sum_{i=1}^p (y_i - \beta_i)^2 + \lambda_1 \sum_{i=1}^p |\beta_i| + \lambda_2 \sum_{i=1}^{p-1} |\beta_{i+1} - \beta_i|.$$

- First term leads to sparse solutions
- Second term leads to piecewise constant solutions



# Fused lasso for supervised classification (Rapaport et al., 2008)

$$\min_{\beta \in \mathbb{R}^p} \sum_{i=1}^n \ell\left(y_i, \beta^\top x_i\right) + \lambda_1 \sum_{i=1}^p |\beta_i| + \lambda_2 \sum_{i=1}^{p-1} |\beta_{i+1} - \beta_i|.$$

where  $\ell$  is, e.g., the hinge loss  $\ell(y, t) = max(1 - yt, 0)$ .

#### Implementation

- When  $\ell$  is the hinge loss (fused SVM), this is a linear program -> up to  $p=10^3\sim 10^4$
- When  $\ell$  is convex and smooth (logistic, quadratic), efficient implementation with proximal methods -> up to  $p=10^8\sim 10^9$

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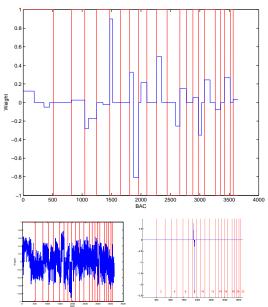
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# Example: predicting metastasis in melanoma

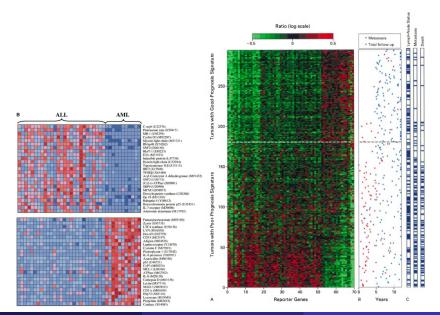


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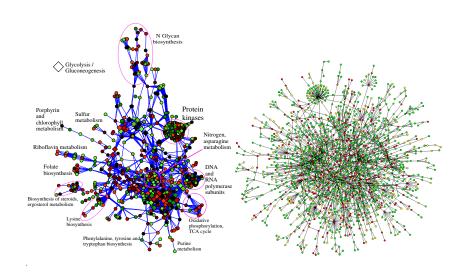
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# Molecular diagnosis / prognosis / theragnosis



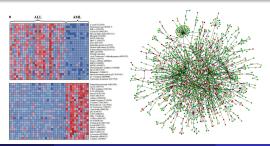
## Gene networks



# Gene networks and expression data

#### Motivation

- Basic biological functions usually involve the coordinated action of several proteins:
  - Formation of protein complexes
  - Activation of metabolic, signalling or regulatory pathways
- Many pathways and protein-protein interactions are already known
- Hypothesis: the weights of the classifier should be "coherent" with respect to this prior knowledge



# Graph-based penalty

$$\min_{\beta} R(\beta) + \lambda \Omega_G(\beta)$$

## Hypothesis

We would like to design penalties  $\Omega_G(\beta)$  to promote one of the following hypothesis:

- Hypothesis 1: genes near each other on the graph should have similar weights (but we do not try to select only a few genes), i.e., the classifier should be smooth on the graph
- Hypothesis 2: genes selected in the signature should be connected to each other, or be in a few known functional groups, without necessarily having similar weights.

# Graph based penalty

#### Prior hypothesis

Genes near each other on the graph should have similar weigths.

#### An idea (Rapaport et al., 2007)

$$\Omega_{spectral}(\beta) = \sum_{i \sim i} (\beta_i - \beta_j)^2$$

$$\min_{eta \in \mathbb{R}^p} R(eta) + \lambda \sum_{i \sim j} (eta_i - eta_j)^2$$

# Graph based penalty

#### Prior hypothesis

Genes near each other on the graph should have similar weigths.

#### An idea (Rapaport et al., 2007)

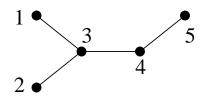
$$\Omega_{spectral}(\beta) = \sum_{i \sim j} (\beta_i - \beta_j)^2 \,,$$

$$\min_{\beta \in \mathbb{R}^p} R(\beta) + \lambda \sum_{i \sim j} (\beta_i - \beta_j)^2$$
.

# Graph Laplacian

#### **Definition**

The Laplacian of the graph is the matrix L = D - A.



$$L = D - A = \begin{pmatrix} 1 & 0 & -1 & 0 & 0 \\ 0 & 1 & -1 & 0 & 0 \\ -1 & -1 & 3 & -1 & 0 \\ 0 & 0 & -1 & 2 & -1 \\ 0 & 0 & 0 & 1 & 1 \end{pmatrix}$$

# Spectral penalty as a kernel

#### **Theorem**

The function  $f(x) = \beta^{\top} x$  where b is solution of

$$\min_{\beta \in \mathbb{R}^p} \frac{1}{n} \sum_{i=1}^n I\left(\beta^\top x_i, y_i\right) + \lambda \sum_{i \sim j} \left(\beta_i - \beta_j\right)^2$$

is equal to  $g(x) = \gamma^{\top} \Phi(x)$  where  $\gamma$  is solution of

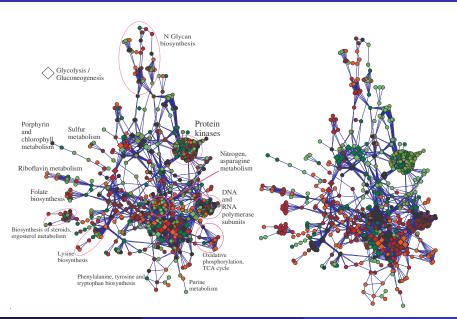
$$\min_{\gamma \in \mathbb{R}^p} \frac{1}{n} \sum_{i=1}^n I\left(\gamma^{\top} \Phi(\mathbf{x}_i), \mathbf{y}_i\right) + \lambda \gamma^{\top} \gamma,$$

and where

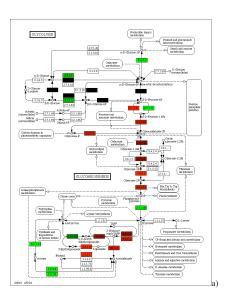
$$\Phi(x)^{\top}\Phi(x') = x^{\top}K_Gx'$$

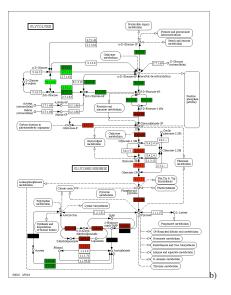
for  $K_G = L^*$ , the pseudo-inverse of the graph Laplacian.

#### Classifiers



#### Classifier





# Other penalties with kernels

$$\Phi(x)^{\top}\Phi(x') = x^{\top}K_Gx'$$

with:

•  $K_G = (c + L)^{-1}$  leads to

$$\Omega(\beta) = c \sum_{i=1}^{p} \beta_i^2 + \sum_{i \sim j} (\beta_i - \beta_j)^2.$$

• The diffusion kernel:

$$K_G = \exp_M(-2tL)$$
.

penalizes high frequencies of  $\beta$  in the Fourier domain.

## Other penalties without kernels

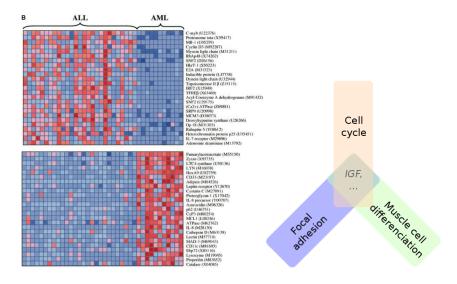
Gene selection + Piecewise constant on the graph

$$\Omega(\beta) = \sum_{i \sim j} |\beta_i - \beta_j| + \sum_{i=1}^p |\beta_i|$$

• Gene selection + smooth on the graph

$$\Omega(\beta) = \sum_{i \sim j} (\beta_i - \beta_j)^2 + \sum_{i=1}^p |\beta_i|$$

# How to select jointly genes belonging to predefined pathways?



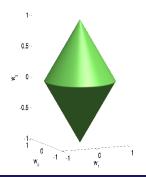
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# Selecting pre-defined groups of variables

## Group lasso (Yuan & Lin, 2006)

If groups of covariates are likely to be selected together, the  $\ell_1/\ell_2$ -norm induces sparse solutions at the group level:

$$\Omega_{group}(w) = \sum_{g} \|w_g\|_2$$



$$\Omega(\mathbf{w}_1, \mathbf{w}_2, \mathbf{w}_3) = \|(\mathbf{w}_1, \mathbf{w}_2)\|_2 + \|\mathbf{w}_3\|_2$$

# What if a gene belongs to several groups?

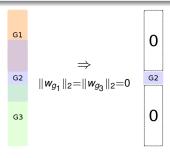
## Issue of using the group-lasso

- $\Omega_{group}(w) = \sum_{g} \|w_g\|_2$  sets groups to 0.
- One variable is selected 

  all the groups to which it belongs are selected.



IGF selection ⇒ selection of unwanted groups



Removal of *any* group containing a gene ⇒ the weight of the gene is 0.

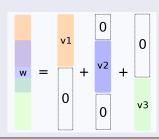
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# Overlap norm (Jacob et al., 2009)

#### An idea

Introduce latent variables  $v_g$ :

$$\left\{egin{aligned} \min_{w,v} L(w) + \lambda \sum_{g \in \mathcal{G}} \|v_g\|_2 \ w = \sum_{g \in \mathcal{G}} v_g \ \mathrm{supp}\left(v_g
ight) \subseteq g. \end{aligned}
ight.$$



#### **Properties**

- Resulting support is a union of groups in G.
- Possible to select one variable without selecting all the groups containing it.
- Equivalent to group lasso when there is no overlap

#### A new norm

## Overlap norm

$$\begin{cases} \min_{w,v} L(w) + \lambda \sum_{g \in \mathcal{G}} \|v_g\|_2 \\ w = \sum_{g \in \mathcal{G}} v_g \\ \text{supp } (v_g) \subseteq g. \end{cases} = \min_{w} L(w) + \lambda \Omega_{\textit{overlap}}(w)$$

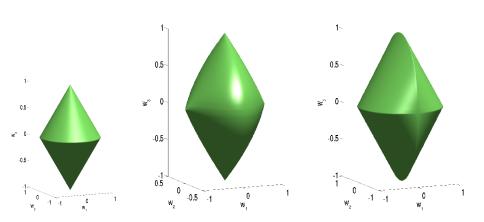
with

$$egin{aligned} \Omega_{\mathit{overlap}}(w) & riangleq egin{aligned} \min_{v} \sum_{g \in \mathcal{G}} \|v_g\|_2 \ w &= \sum_{g \in \mathcal{G}} v_g \ \mathrm{supp}\left(v_a
ight) \subseteq g. \end{aligned}$$

#### **Property**

- $\Omega_{overlap}(w)$  is a norm of w.
- $\Omega_{overlap}(.)$  associates to w a specific (not necessarily unique) decomposition  $(v_g)_{g \in \mathcal{G}}$  which is the argmin of (\*).

# Overlap and group unity balls



Balls for  $\Omega^{\mathcal{G}}_{\mathsf{group}}(\cdot)$  (middle) and  $\Omega^{\mathcal{G}}_{\mathsf{overlap}}(\cdot)$  (right) for the groups  $\mathcal{G} = \{\{1,2\},\{2,3\}\}$  where  $w_2$  is represented as the vertical coordinate. Left: group-lasso  $(\mathcal{G} = \{\{1,2\},\{3\}\})$ , for comparison.

## Theoretical results

## Consistency in group support (Jacob et al., 2009)

- Let  $\bar{w}$  be the true parameter vector.
- Assume that there exists a unique decomposition  $\bar{v}_g$  such that  $\bar{w} = \sum_g \bar{v}_g$  and  $\Omega_{\text{overlap}}^{\mathcal{G}}\left(\bar{w}\right) = \sum \|\bar{v}_g\|_2$ .
- Consider the regularized empirical risk minimization problem  $L(w) + \lambda \Omega_{\text{overlap}}^{\mathcal{G}}(w)$ .

#### Then

- under appropriate mutual incoherence conditions on *X*,
- as  $n \to \infty$ ,
- with very high probability,

the optimal solution  $\hat{w}$  admits a unique decomposition  $(\hat{v}_g)_{g \in \mathcal{G}}$  such that

$$ig\{g\in\mathcal{G}|\hat{v}_g
eq0ig\}=ig\{g\in\mathcal{G}|ar{v}_g
eq0ig\}$$
 .

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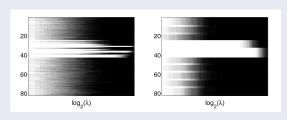
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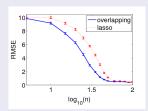
$$\left\{g\in\mathcal{G}|\hat{v}_g
eq 0
ight\}=\left\{g\in\mathcal{G}|ar{v}_g
eq 0
ight\}.$$

# Experiments

#### Synthetic data: overlapping groups

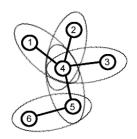
- 10 groups of 10 variables with 2 variables of overlap between two successive groups :{1,...,10}, {9,...,18},...,{73,...,82}.
- Support: union of 4th and 5th groups.
- Learn from 100 training points.





Frequency of selection of each variable with the lasso (left) and  $\Omega_{\text{overlap}}^{\mathcal{G}}(.)$  (middle), comparison of the RMSE of both methods (right).

# Graph lasso



#### Two solutions

$$\Omega_{\textit{intersection}}(\beta) = \sum_{i \sim j} \sqrt{\beta_i^2 + \beta_j^2} \,,$$

$$\Omega_{\textit{union}}(\beta) = \sup_{\alpha \in \mathbb{R}^p: \forall i \sim j, \|\alpha_j^2 + \alpha_j^2\| \leq 1} \alpha^\top \beta \ .$$

# Graph lasso vs kernel on graph

• Graph lasso:

$$\Omega_{ ext{graph lasso}}( extbf{ extit{w}}) = \sum_{i \sim j} \sqrt{ extbf{ extit{w}}_i^2 + extbf{ extit{w}}_j^2} \,.$$

constrains the sparsity, not the values

Graph kernel

$$\Omega_{ ext{graph kernel}}(w) = \sum_{i \sim j} (w_i - w_j)^2$$
 .

constrains the values (smoothness), not the sparsity

# Preliminary results

#### Breast cancer data

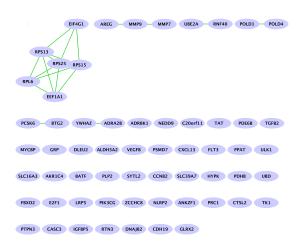
- Gene expression data for 8, 141 genes in 295 breast cancer tumors.
- Canonical pathways from MSigDB containing 639 groups of genes, 637 of which involve genes from our study.

METHOD	$\ell_1$	$\Omega_{OVERLAP}^{\mathcal{G}}\left(.\right)$
ERROR	$\textbf{0.38} \pm \textbf{0.04}$	$\textbf{0.36} \pm \textbf{0.03}$
MEAN ♯ PATH.	130	30

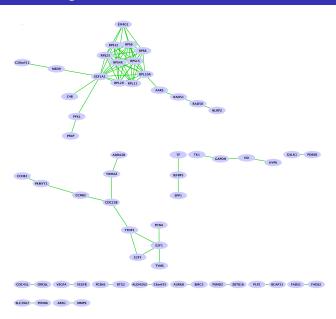
Graph on the genes.

METHOD	$\ell_1$	$\Omega_{graph}(.)$
ERROR	$\textbf{0.39} \pm \textbf{0.04}$	$\textbf{0.36} \pm \textbf{0.01}$
Av. SIZE C.C.	1.03	1.30

## Lasso signature



## Graph Lasso signature



## **Outline**

- Motivations
- 2 Finding multiple change-points in a single profile
- Finding multiple change-points shared by many signals
- Supervised classification of genomic profiles
- 5 Learning molecular classifiers with network information
- 6 Conclusion

#### Conclusions

- Feature / pattern selection in high dimension is central for many applications
- Convex sparsity-inducing penalties or positive definite kernels are promising
- Success stories remain limited on real data...
- Need to adjust the complexity of the model to the data available

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